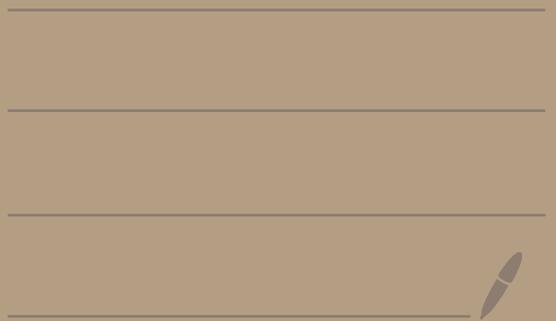


Math 4740

4/21/25



Topic 8 - Continuous random variables

Def: Let $f: \mathbb{R} \rightarrow \mathbb{R}$. We say that f is a probability density function (pdf) if

① $f(x) \geq 0$ for all x

② $\int_{-\infty}^{\infty} f(x) dx$ exists and $\int_{-\infty}^{\infty} f(x) dx = 1$



Ex: (Standard normal distribution)

$$\text{Let } f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$$

Let's show that f is a pdf.

$$\textcircled{1} f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2} > 0 \quad \text{for all } x$$

$$\textcircled{2} \text{ Let } I = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx.$$

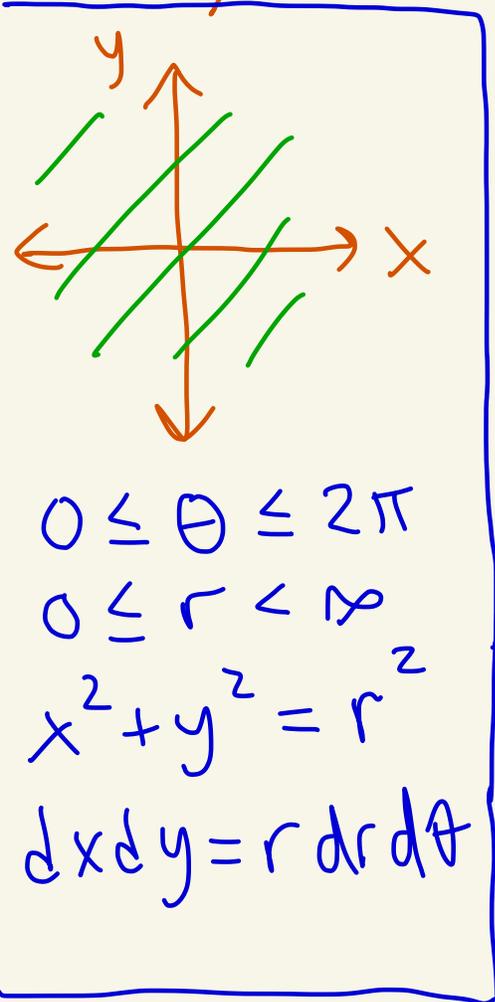
Then,

$$I^2 = \left(\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx \right) \left(\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-y^2/2} dy \right)$$

$$= \int_{-\infty}^{\infty} \left(\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-y^2/2} dy \right) \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-x^2/2} e^{-y^2/2} dx dy$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-(x^2+y^2)/2} dx dy$$



$$= \frac{1}{2\pi} \int_0^{2\pi} \int_0^{\infty} e^{-r^2/2} r dr d\theta$$

$$= \frac{1}{2\pi} \int_0^{2\pi} \left[-e^{-r^2/2} \Big|_0^{\infty} \right] d\theta$$

$$= \frac{1}{2\pi} \int_0^{2\pi} \left[0 - \left(-e^{-0^2/2} \right) \right] d\theta$$

$$= \frac{1}{2\pi} \int_0^{2\pi} 1 d\theta = \frac{1}{2\pi} \theta \Big|_0^{2\pi} = \frac{1}{2\pi} (2\pi - 0) = 1$$

So, $I^2 = 1$. Since $I \geq 0$

We know $I = 1$.

Thus, $\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = 1$.

So, f is a pdf.

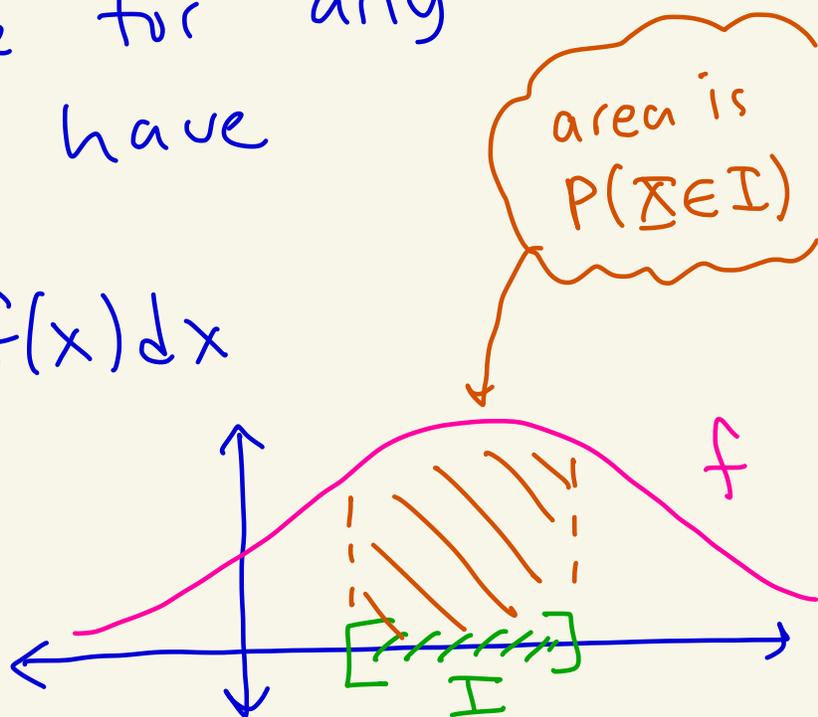
Def: Let X be a random variable.

We say that X is a continuous
random variable if there exists

a pdf f where for any
interval I we have

$$P(X \in I) = \int_I f(x) dx$$

For example,



$$P(a \leq X \leq b) = \int_a^b f(x) dx$$

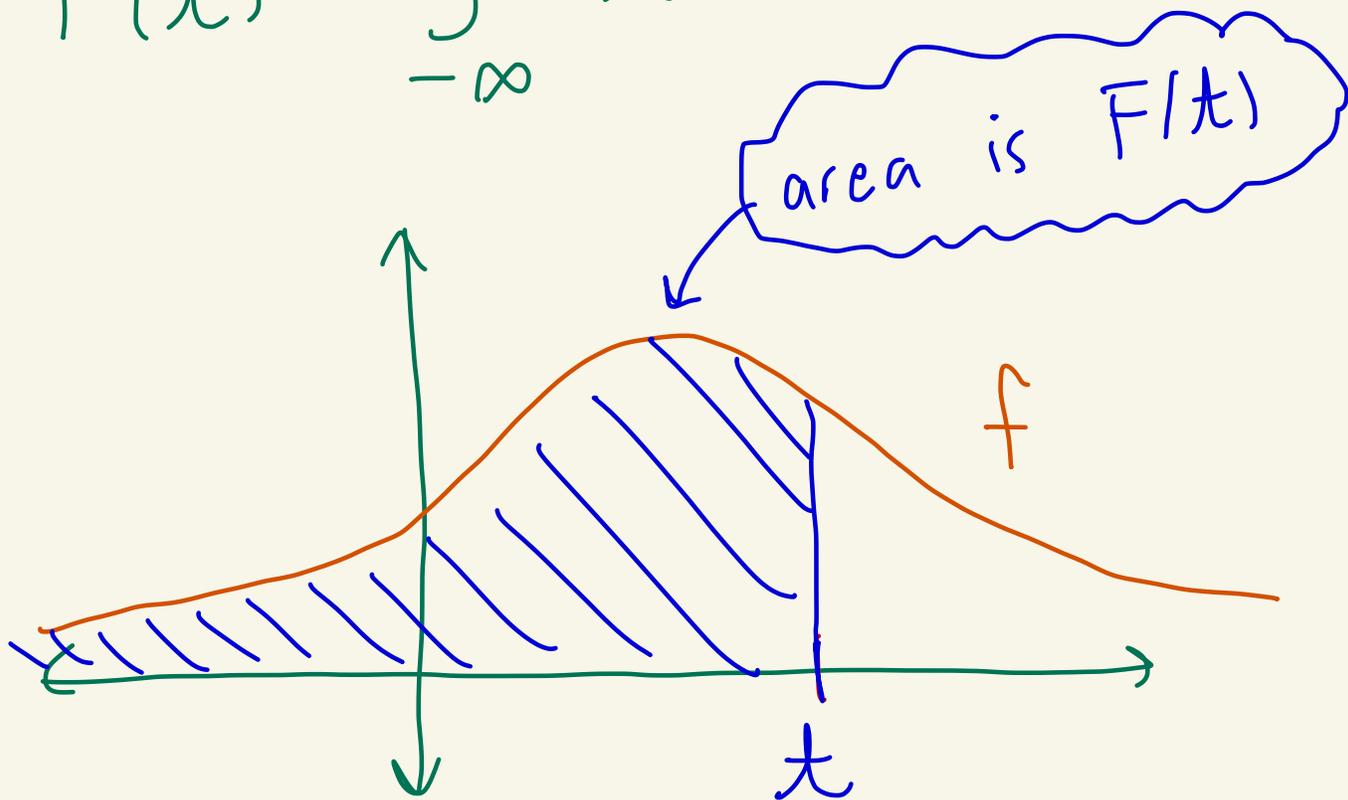
$$P(a \leq \bar{X}) = \int_a^{\infty} f(x) dx$$

$$P(X \leq b) = \int_{-\infty}^b f(x) dx$$

The cumulative distribution function

(cdf) of X is

$$F(t) = \int_{-\infty}^t f(x) dx$$

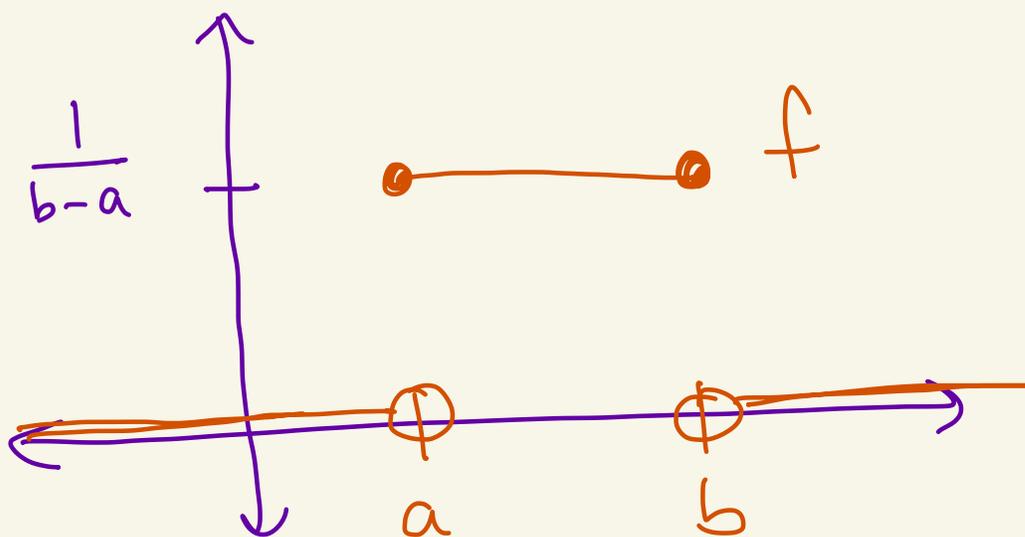


Note by the fundamental theorem of calculus, F is an anti-derivative of f , that is $F' = f$.

Ex: (the uniform distribution on $[a, b]$)

Let $a < b$.

$$f(x) = \begin{cases} \frac{1}{b-a} & \text{if } a \leq x \leq b \\ 0 & \text{otherwise} \end{cases}$$

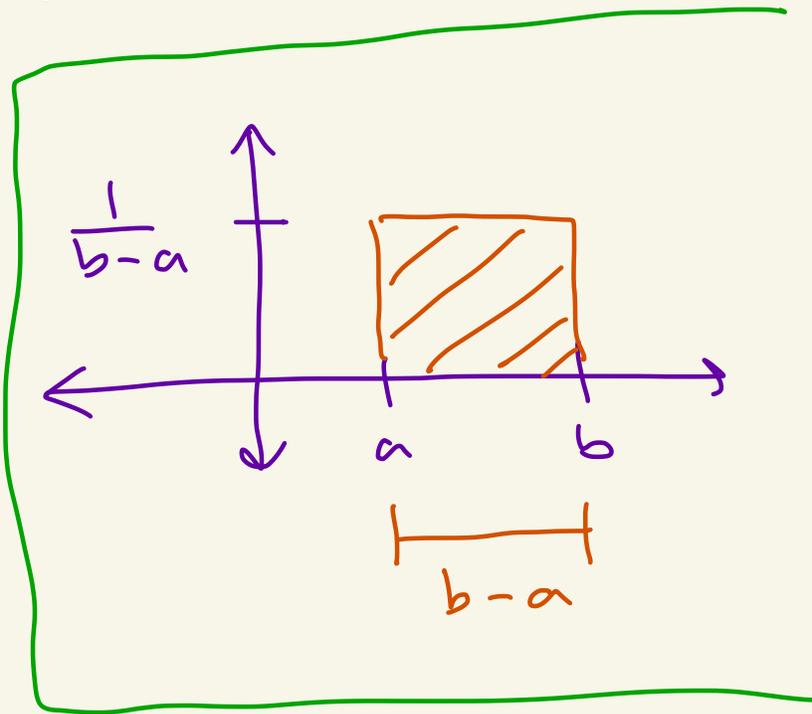


f is a pdf:

① $f(x) \geq 0$

② $\int_{-\infty}^{\infty} f(x) dx = \int_a^b \frac{1}{b-a} dx$

$$= \underbrace{\left(\frac{1}{b-a}\right)}_{\text{height}} \underbrace{(b-a)}_{\text{base}} = 1$$

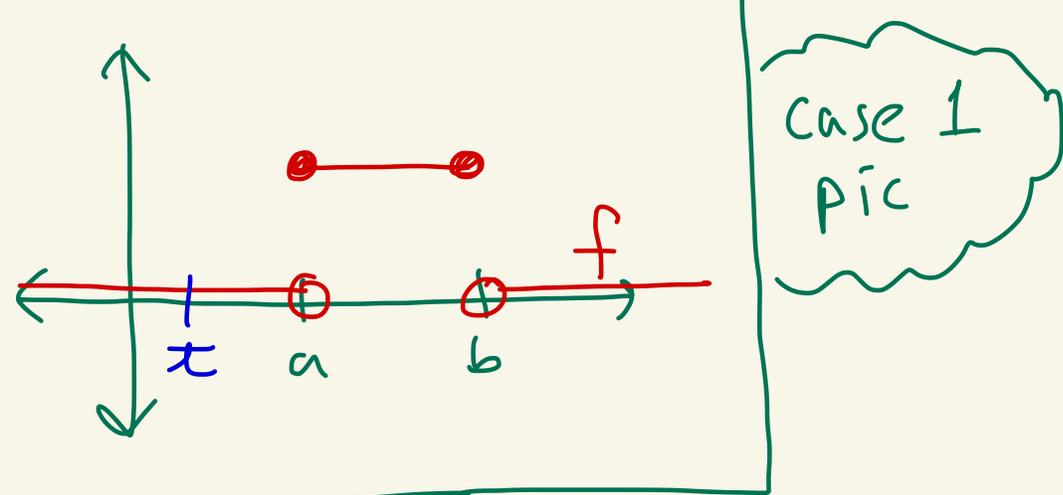


Let $F(t) = \int_{-\infty}^t f(x) dx$

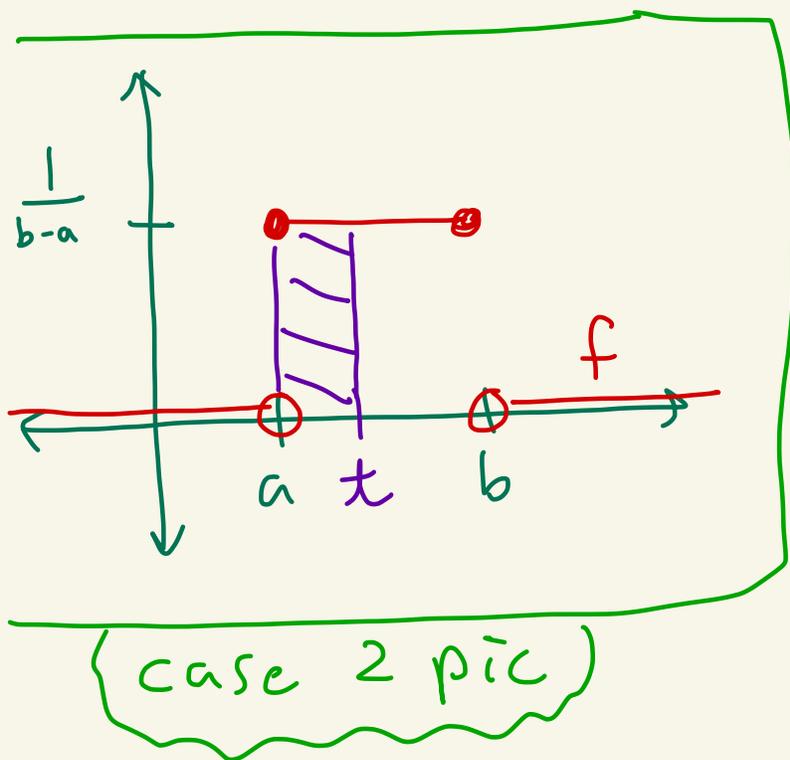
Let's find a formula for F.

case 1: Suppose $t \leq a$. Then,

$$F(t) = \int_{-\infty}^t f(x) dx = \int_{-\infty}^t 0 dx = 0.$$

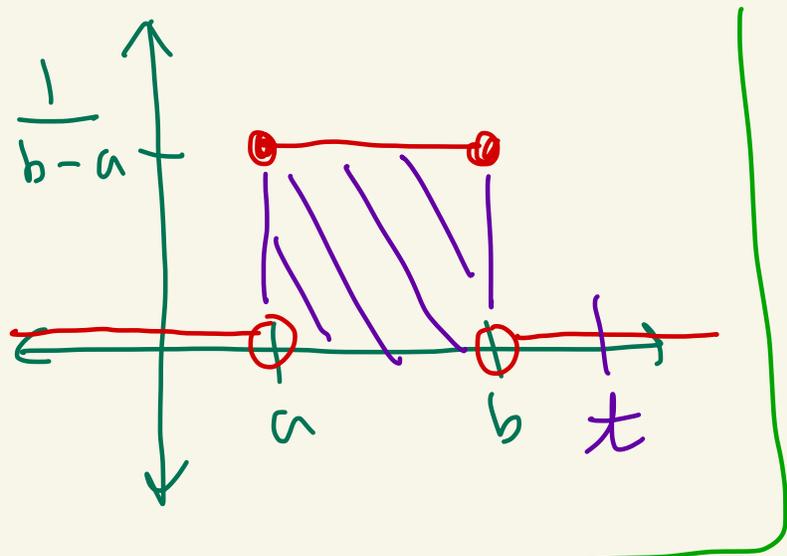


Case 2: Suppose $a \leq t \leq b$.



$$\begin{aligned}
 F(t) &= \int_{-\infty}^t f(x) dx \\
 &= \int_a^t \frac{1}{b-a} dx \\
 &= \underbrace{\left(\frac{1}{b-a}\right)}_{\text{base}} \underbrace{(t-a)}_{\text{height}} \\
 &= \frac{t-a}{b-a}
 \end{aligned}$$

Case 3: Suppose $b \leq t$.

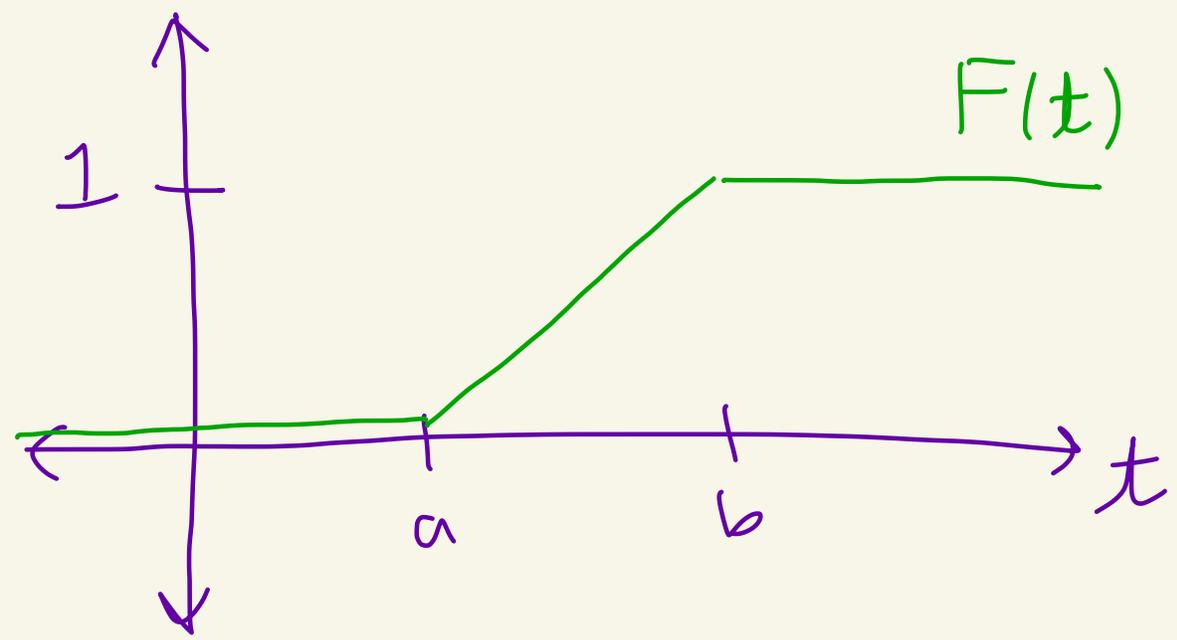


$$F(t) = \int_{-\infty}^t f(x) dx = 1$$

Case 3

Thus,

$$F(t) = \begin{cases} 0 & \text{if } t \leq a \\ \frac{t-a}{b-a} & \text{if } a \leq t \leq b \\ 1 & \text{if } b \leq t \end{cases}$$

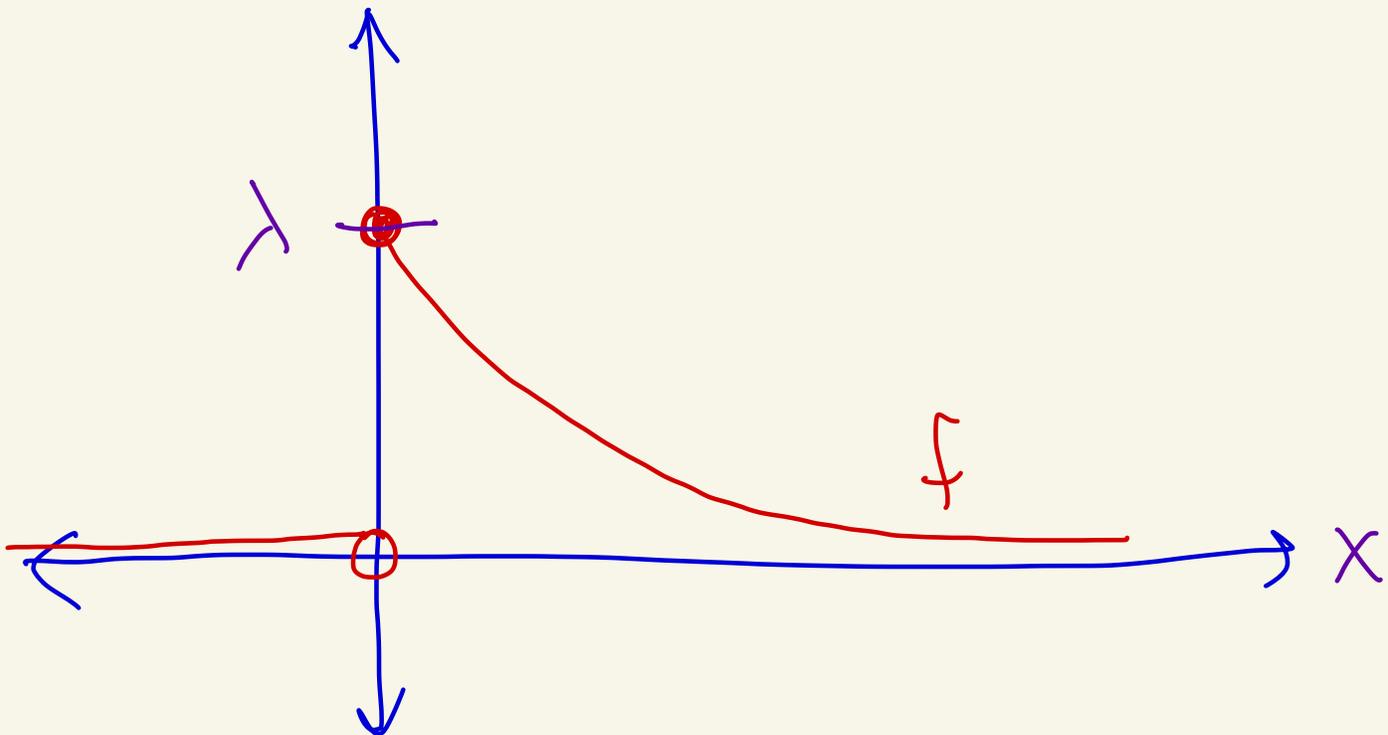


Ex: (Exponential random variable
with parameter λ)

Let $\lambda > 0$.

Define

$$f(x) = \begin{cases} 0 & \text{if } x < 0 \\ \lambda e^{-\lambda x} & \text{if } x \geq 0 \end{cases}$$



Let's show that f is a pdf.

① $f(x) \geq 0$

② $\int_{-\infty}^{\infty} f(x) dx = \int_0^{\infty} \lambda e^{-\lambda x} dx$

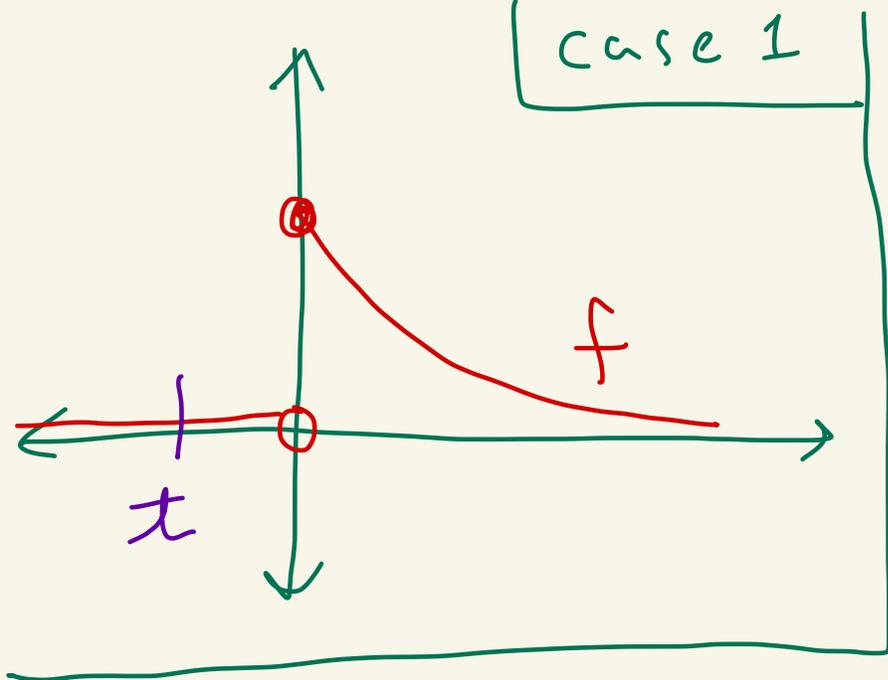
$$= -e^{-\lambda x} \Big|_0^{\infty} = \left[\underset{\substack{\uparrow \\ x \rightarrow \infty}}{-0} - (-e^0) \right]$$

$$= 0 - (-1) = 1$$

Let's find the cumulative distribution function.

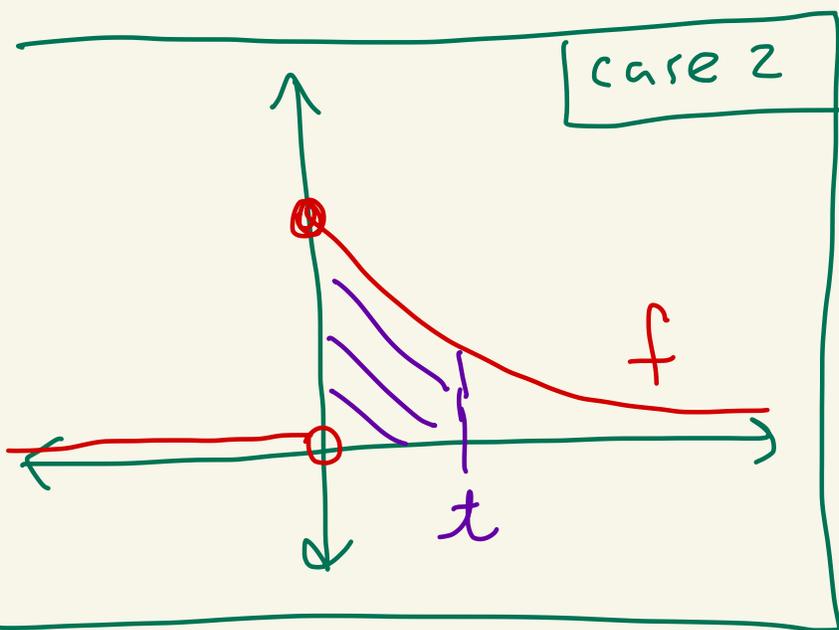
Case 1: Suppose $t < 0$. Then,

$$F(t) = \int_{-\infty}^t f(x) dx = \int_{-\infty}^t 0 dx = 0$$



case 2: Suppose $t \geq 0$. Then,

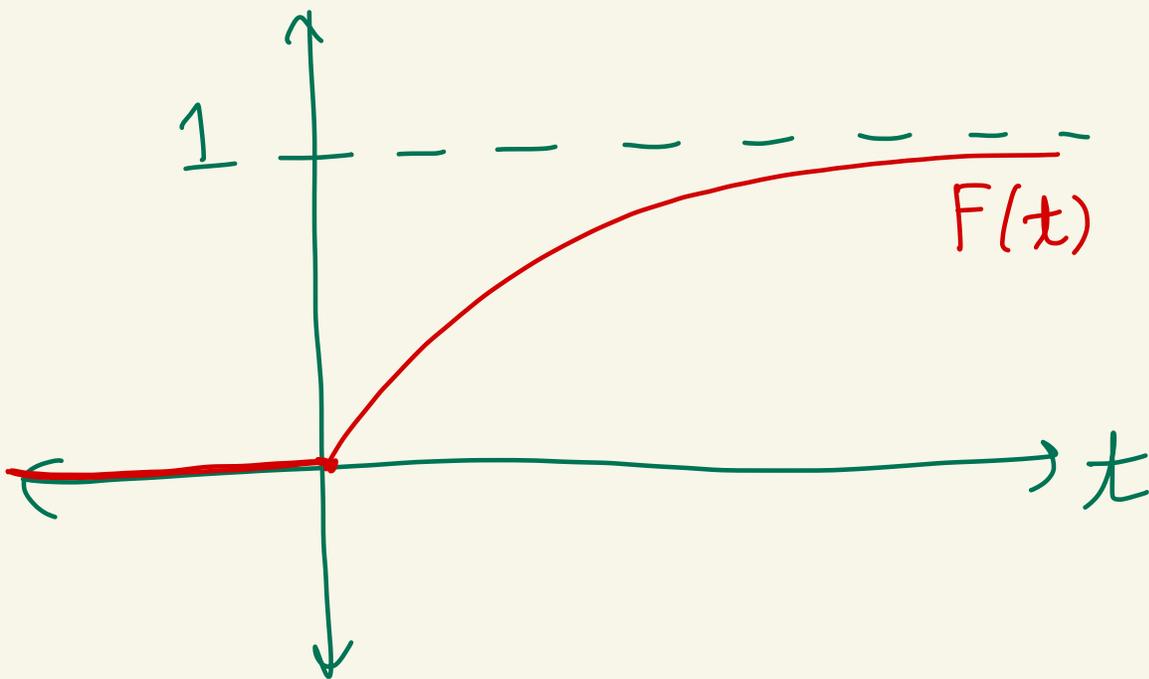
$$F(t) = \int_{-\infty}^t f(x) dx = \int_0^t \lambda e^{-\lambda x} dx$$



$$\begin{aligned}
 &= -e^{-\lambda x} \Big|_0^t \\
 &= -e^{-\lambda t} - (-e^0) \\
 &= -e^{-\lambda t} + 1 \\
 &= 1 - e^{-\lambda t}
 \end{aligned}$$

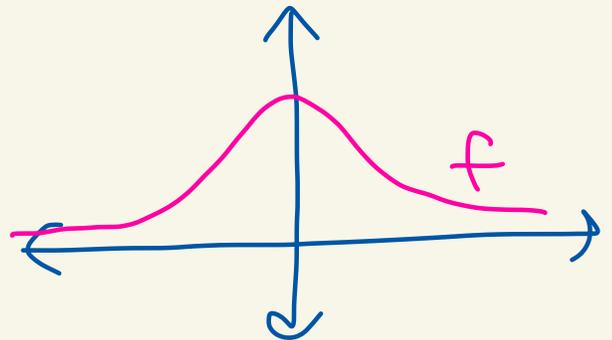
So,

$$F(t) = \begin{cases} 0 & \text{if } t < 0 \\ 1 - e^{-\lambda t} & \text{if } 0 \leq t \end{cases}$$



Ex: (Standard normal distribution)

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$$



$$F(t) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^t e^{-x^2/2} dx = \underbrace{\Phi(t)}_{\text{topic 7}}$$

Def: Let \bar{X} be a continuous random variable with pdf f .

Then the expected value of \bar{X} is

$$E[\bar{X}] = \int_{-\infty}^{\infty} x f(x) dx$$

Let $\mu = E[\bar{X}]$.

The variance is

$$\text{Var}(\bar{X}) = \left(\int_{-\infty}^{\infty} x^2 f(x) dx \right) - \mu^2$$

and the standard deviation is

$$\sigma_{\bar{X}} = \sqrt{\text{Var}(\bar{X})}$$